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Capacity of DNA Data Embedding Under Substitution Mutations
Félix Balado, Member, IEEE

Abstract—A number of methods have been proposed over the last decade for encoding information using deoxyribonucleic acid (DNA), giving rise to the emerging area of DNA data embedding. Since a DNA sequence is conceptually equivalent to a sequence of quaternary symbols (bases), DNA data embedding (diversely called DNA watermarking or DNA steganography) can be seen as a digital communications problem where channel errors are analogous to mutations of DNA bases. Depending on the use of coding or noncoding DNA host sequences, which respectively denote DNA segments that can or cannot be translated into proteins, DNA data embedding is essentially a problem of communications with or without side information at the encoder. In this paper the Shannon capacity of DNA data embedding is obtained for the case in which DNA sequences are subject to substitution mutations modelled using the Kimura model from molecular evolution studies. Inferences are also drawn with respect to the biological implications of some of the results presented.

Index Terms—Channel capacity, DNA, Genetic Communication, Watermarking, Evolution (biology)

I. INTRODUCTION

The last ten years have seen the proposal of numerous practical methods [1], [2], [3], [4], [5], [6], [7], [8], [9], [10], [11], [12] for encoding nongenetic information using DNA molecules as a medium, both in vitro and in vivo. A high profile use of these techniques took place recently, when the J. Craig Venter Institute (JCVI) produced the first artificial bacteria which included a “watermark” containing authorship information [11], [12]. All of these proposals rely on the fact that DNA molecules—which encode genetic information in all living organisms, except for some viruses—are conceptually equivalent to sequences of quaternary symbols, which can be manipulated to store and retrieve arbitrary data by using the molecular biology analogues of “writing” (producing recombinant DNA) and “reading” (DNA sequencing). Therefore DNA data embedding is in essence an instance of digital communications, with the particularity that channel errors are analogous to mutations of DNA components. The two broad fields of application of DNA data embedding techniques are: 1) the use of DNA strands as self-replicating nano-memories with the ability to store huge amounts of data in an ultra-compact and energy-efficient way [2], [4], [7], [8]; and 2) security and tracking applications made possible by embedding nongenetic information in DNA: for instance, DNA watermarking [9], [11], [12], DNA steganography [1], [5] and DNA tagging [13]. The actual information embedded in security applications may include ownership, authorship and integrity data.

DNA data embedding techniques can be divided into two types, depending on whether the embedded information is hosted by coding or by noncoding parts of a genome—the ensemble of all the DNA of an organism. The coding parts of a genome are the genes that it contains, which are the basic units of genetic inheritance. The remaining parts of a genome are generically called noncoding DNA. Noncoding DNA is normally the host of choice in storage applications of DNA data embedding: it will be seen that more information can usually be embedded in noncoding sections, with less constraints and complications. However it is also important to consider genes (that is, coding DNA) as information hosts for several reasons. Firstly, embedding information within genes may be the only approach to mark or track them in a truly individualised way, which is of relevance in some security and tracking applications. One example is the potential use of DNA data embedding in gene patenting [5], [14], a procedure by which intellectual property of artificially engineered genes may be asserted. Patent information must reside within a gene—and not anywhere else—if that information is to travel with the gene when copied by an unauthorised party. It may also be desirable to include laboratory identification or integrity information in genes employed in gene therapies (human gene transfer), in which case the information must lie within genes [3]. Furthermore, Arita and Ohashi point out that it may be biologically safer to embed information within coding DNA rather than within noncoding DNA [5]. This may seem paradoxical, because noncoding DNA was once derogatorily labelled as “junk DNA” due to its supposed lack of biological purpose. However it is now increasingly recognised and understood that specific noncoding DNA sections can be involved in regulatory functions of gene activity. Therefore risks may lie ahead if information is embedded in noncoding regions which may hold as yet unknown functions.

DNA data embedding methods which target the genomes of self-replicating living organisms—such as the ones proposed and implemented by Wong et al. [4], Arita and Ohashi [5], Yachie et al. [7], Heider and Barnekow [14], and more recently Gibson et al. [12]—all face two fundamental questions: 1) how much information can be embedded in a given DNA segment of an organism?; and 2) how many generations of that organism can this information potentially survive intact? The
methods just mentioned embed information at rates ranging from 2 bits per DNA base [4], [12] (using noncoding DNA) to 0.2136 bits per base [14] and 0.0926 bits per base [5] (using coding DNA), but will that information be retrievable without errors when mutations accumulate after a number of generations of the host organism? Could those rates possibly be higher for a given mutation rate? The answers are supplied by the upper limit on the amount of information that can be embedded within a given DNA molecule under a given mutation rate, and asymptotically decoded with no errors. This is the Shannon capacity [15] of DNA data embedding, which is the ultimate benchmark against which all actual DNA data embedding methods used in bioengineering applications must be measured.

In this paper the capacity of DNA data embedding under substitution mutations is determined. Substitution mutations are those that randomly switch bases in a DNA sequence, as may occur in the replication of an organism. These are modelled by means of a symmetric memoryless channel equivalent to the probabilistic model first proposed by Kimura [16] to study molecular evolution. Section II describes the framework, assumptions and model used. The capacity analysis, addressed in Section III, is straightforward when no side information is required by the encoder during the embedding process. This is the case for techniques that use noncoding DNA to host information, such as [4], [7] or [12]. On the other hand, the capacity analysis with side information at the encoder corresponds to techniques which use coding DNA, such as [3], [5], [9] or [14], whose encoder must strictly preserve gene expression. This analysis forms the main constituent of this paper. Where pertinent, biological implications of these information theoretical results are also discussed. In particular, the non side-informed scenario happens to be closely connected to previous studies by May [17], Battail [18], and other authors who have applied information theoretical concepts to molecular biology from an information transmission perspective.

II. PRELIMINARY CONCEPTS AND ASSUMPTIONS

Chemically, DNA is formed by two backbone strands helically twisted around each other, and mutually attached by means of two nitrogenous base sequences. The four possible bases are the molecules adenine, cytosine, thymine and guanine, abbreviated to A, C, T and G, respectively. Only the pairings A–T and C–G can exist between the two strands, which is why each of the two base sequences is completely determined by the other, and also why the length of a DNA molecule is measured in base pairs. The interpretation of DNA as a one-dimensional discrete digital signal is straightforward: any one of the two strands constitutes a digital sequence formed by symbols from a quaternary alphabet, which completely determines the DNA molecule.

Codons —the minimal biological “codewords”— are formed by triplets of consecutive bases in a base sequence. Given any three consecutive bases there is no ambiguity in the codon they stand for, since there is only one direction in which a base sequence can be read. In molecular biology this is called the 5′–3′ direction, which refers to certain chemical feature points in a DNA backbone strand. The two strands in a DNA molecule are read in opposite directions; because of this, and because of their complementarity, they are termed antiparallel. Groups of consecutive codons in some special regions of a DNA sequence can be translated into a series of chemical compounds called amino acids via transcription to the intermediary ribonucleic acid (RNA) molecule. RNA is similar to DNA but single-stranded and with uracil (abbreviated U) replacing thymine. Amino acids are sequentially assembled in the same order imposed by the codon sequence. The result of this assembling process are proteins, which are the basic compounds of the chemistry of life. There are 4^3 = 64 possible codons, since they are triplets of 4-ary symbols. Crucially, there are only 20 possible amino acids, mapped to the 64 codons according to the almost universal genetic code. The standard version\(^1\) of the genetic code is shown in Table I and explained in more detail later.

The genome of an organism is the ensemble of all its DNA. Segments of a genome that can be translated into proteins by the genetic machinery through the process described above are called protein-coding DNA (pcDNA), or just coding DNA, whereas those segments that never get translated are called noncoding DNA (ncDNA). A gene is a pcDNA segment, or group of segments, which encodes one single protein, and which is flanked by certain start and stop codons (see Table I) plus other regulatory markers. Finally, for each base sequence there are three different reading frames which determine three different codon sequences. The correct reading frame is marked by the position of a start codon.

The main assumptions made in this work are:

- **ncDNA can be freely appended or overwritten.** As briefly mentioned in the introduction this assumption is not always strictly true, because certain ncDNA regions act as promoters or enhancers for gene expression, or are transcribed into regulatory RNA (which is not translated into proteins). After conducting in vivo tests, Heider et al. have cautioned against embedding information in promoter ncDNA regions [10]. However the working hypothesis is valid in many other suitably chosen ncDNA regions, as proved by Wong et al. [4], Yachie et al. [7], [8] and Gibson et al. [12], who have successfully embedded information in the ncDNA of living organisms.

- **pcDNA can be freely modified as long as the genetic code is followed.** This is the classic standard assumption supporting the validity of the genetic code. This assumption was used in silico by Shimanovsky et al. [3] and in vivo by Arita and Ohashi [5] and Heider and Barnekow [14]. However, living organisms also exhibit preferred codon usage biases (also called codon usage statistics), which are characteristic distributions of the codons associated with a given amino acid. When pcDNA is modified, these codon biases can emerge completely changed even if the genetic code is strictly observed. This change might be detrimental for gene expression, for instance by extending gene translation times among other effects [19]. For this reason we will also analyse the embedding capacity

\(^1\)http://www.ncbi.nlm.nih.gov/taxonomy (genetic codes)
of pcDNA with codon bias preservation constraints. To conclude, it is important to mention that only nonoverlapping genes will be considered, either on the same or on opposite strands. In any case overlapping genes are rare occurrences, except in very compact genomes.

**Notation.** Calligraphic letters ($X$) denote sets; $|X|$ is the cardinality of $X$. Boldface letters ($x$) denote row vectors, and $I$ is an all-ones vector. If a Roman letter is used both in uppercase ($X$) and in lowercase ($x$), the two forms denote a random variable and a realisation of it, respectively. $p(X = x)$ is the probability mass function (pmf) or distribution of $X$; we will simply write $p(x)$ when the variable is clear from the context. For brevity, $X$ may also denote the distribution of $X$ if there is no ambiguity. $E[X]$ is the mathematical expectation of $X$, and $H(X)$ its entropy. Also, $h(q)$ is the entropy of a Bernoulli($q$) random variable. $I(X; Y)$ is the mutual information between $X$ and $Y$. Logarithms are base 2, unless explicitly stated otherwise. The Hamming distance between vectors $x$ and $y$ is denoted by $d_H(x, y)$.

A ncDNA sequence will be denoted by a vector $\mathbf{x} = [x_1, x_2, \ldots, x_n]$, whose elements are consecutive bases from a base sequence. That is, $x_i \in \mathcal{X} \triangleq \{A, C, T, G\}$, the 4-ary set of possible bases. A pcDNA sequence will be denoted by a vector of vectors $\mathbf{X} = [\mathbf{x}_1, \mathbf{x}_2, \ldots, \mathbf{x}_n]$, whose elements are consecutive codons of the two antiparallel base sequences when the right reading frame among the three possibilities is chosen. Therefore, $x_i \in \mathcal{X}^3$. The amino acid into which a codon $x_i$ uniquely translates is denoted by $x'_i \triangleq \alpha(x_i) \in \mathcal{X}'$, where $\mathcal{X}' \triangleq \{\text{Ala}, \text{Arg}, \text{Asn}, \text{Asp}, \text{Cys}, \text{Gln}, \text{Glu}, \text{Gly}, \text{His}, \text{Ile}, \text{Leu}, \text{Lys}, \text{Met}, \text{Phe}, \text{Pro}, \text{Ser}, \text{Thr}, \text{Trp}, \text{Tyr}, \text{Val}, \text{Stp}\}$ is the set of all amino acids defined using the standard three-letter abbreviations of their names. $\mathbf{x}' = \alpha(\mathbf{X}) = [x'_1, x'_2, \ldots, x'_n]$ denotes the unique amino acid sequence established by $\mathbf{X}$, usually called the primary structure of the protein encoded by $\mathbf{X}$.

The subset of synonymous codons associated with amino acid $x' \in \mathcal{X}'$, namely

$$S_{x'} \triangleq \{x \in \mathcal{X}^3 | \alpha(x) = x'\},$$

is established by the genetic code in Table I. The ensemble of stop codons is collected under the special symbol $\text{Stp}$, and thus loosely classed as an “amino acid” for notational convenience. However $\text{Stp}$ just indicates the end of a gene, and does not actually map to any amino acid. Three codons, underlined in Table I, also double as protein start translation commands when they appear at the start of a gene—the two corresponding to Leu only in eukaryotic organisms. The number $|S_{x'}|$ of synonymous codons mapping to $x'$ is called the multiplicity of $x'$. Due to the uniqueness of the mapping from codons to amino acids, observe that $S_{x'} \cap S_{x''} = \emptyset$ for $x' \neq x'' \in \mathcal{X}'$, and that $\sum_{x' \in \mathcal{X}'} |S_{x'}| = |\mathcal{X}'|^3 = 64$ since $\cup_{x' \in \mathcal{X}'} S_{x'} = \mathcal{X}'^3$. Notice that, for most amino acids $x'$, the third base (wobble base) tends to exhibit more variability than the other two bases in synonymous codons from $S_{x'}$. Finally, a toy example of a pcDNA sequence may be $\mathbf{x} = [\text{Tyr}, \text{Cys}]$, which would encode the amino acid sequence $x' = \alpha(\mathbf{x}) = [\text{Tyr}, \text{Cys}]$. The corresponding base sequence would be $\mathbf{X} = [\text{T}, \text{A}, \text{T}, \text{T}, \text{G}, \text{C}]$.

### A. Mutation channel model

As mentioned in the introduction, an information-carrying DNA molecule undergoing mutations can be readily seen as a digital signal undergoing a noisy communications channel, which may be termed “mutation channel” in this context. Herein we will only consider base substitution mutations (also called point mutations), which randomly switch letters within the DNA alphabet. From a communications viewpoint, apart from base substitutions, the most relevant types of mutations are random base insertions and deletions, which cause a synchronisation problem at the decoder. Insertions and deletions are jointly called *indels*, due to the ambiguity that sometimes exists about whether certain differences between DNA sequences are actually due to either type of mutation. The analysis will not consider indel mutations for the following reasons. On the one hand these mutations tend to have deleterious effects when they occur in pcDNA sections of an organism—especially if the number of such mutations is not a multiple of three, as this can derail protein translation—which arguably makes the survival of the organism less likely. Thus the arrival at the decoder of the mutated information-carrying host will also be less likely. Consequently a base substitutions only analysis is deemed a realistic approach in computing the capacity of pcDNA embedding. This is not the case when indels occur in ncDNA sections; however, the fact is that the exact Shannon capacity of a channel with insertions and deletions is currently unknown even in the simplest case of a binary channel [20]. For this reason, dealing with indel mutations is beyond the scope of this paper. In any case, the substitution mutations analysis still yields an upper bound to...
the capacity of ncDNA embedding, with respect to a more general mutation model.

It will be assumed in our analysis that mutations are mutually independent, which is a worst-case scenario in terms of capacity. As a result it is considered that the channel is memoryless and thus accepts a single-letter characterisation. Consequently we will drop vector element subindices whenever this is unambiguous and notionally convenient. The base substitution channel will be modelled by means of the two-parameter Kimura model of nucleotide substitution [16], which has been extensively used in molecular evolution studies. This consists of a 4 × 4 transition probability matrix \( \Pi \triangleq [p(Z = z | Y = y)] \), where \( z, y \in \mathcal{X} \), which presents the following structure:

\[
\Pi = \begin{bmatrix}
1 - q & \frac{q}{2} & \frac{q}{2} & (1 - \frac{3q}{4})q \\
\frac{q}{2} & 1 - q & 1 - q & \frac{q}{2} \\
\frac{q}{2} & \frac{q}{2} & 1 - q & \frac{q}{2} \\
(1 - \frac{2q}{3})q & \frac{q}{2} & \frac{q}{2} & 1 - q
\end{bmatrix}
\]

(1)

From this definition, the probability of base substitution mutation, or base substitution mutation rate, is

\[
q = p(Z \neq y | Y = y) = \sum_{z \neq y} p(Z = z | Y = y), \quad (2)
\]

for any \( y \in \mathcal{X} \), whereas it must hold that \( 0 \leq \gamma \leq 3/2 \) so that row probabilities add up to one. The particular structure of \( \Pi \) aims at reflecting the fact that DNA bases belong to one of two categories according to their chemical structure: purines, \( \mathcal{R} \triangleq \{A, G\} \), or pyrimidines, \( \mathcal{Y} \triangleq \{C, T\} \). Put simply, purines and pyrimidines are both cyclic compounds. Their main difference is that pyrimidines are single-ringed, whereas purines are double-ringed and one of their rings is in fact a pyrimidine.

There are two types of base substitutions corresponding to these categories, which in biological nomenclature are:

- **Base transitions**: those that preserve the category which the base belongs to. In this case the model establishes that \( p(Z = z | Y = y) = (1 - 2\gamma/3)q \) for \( z \neq y \) when either both \( z, y \in \mathcal{R} \) or both \( z, y \in \mathcal{Y} \).
- **Base transversions**: those that switch the base category. In this case the model establishes that \( p(Z = z | Y = y) = (\gamma/3)q \) for \( z \neq y \) when \( z \in \mathcal{Y} \) and \( z \in \mathcal{R} \), or vice versa.

The channel model (1) can incorporate any given transition/transversion ratio \( \varepsilon \) by setting \( \gamma = 3/(2(\varepsilon + 1)) \). Estimates of \( \varepsilon \) given in [21] for the DNA of different organisms range between 0.89 and 18.67, corresponding to \( \gamma \) between 0.07 and 0.79. This range of \( \varepsilon \) reflects the fact that base transitions are generally much more probable than base transversions due to the chemical structure similarity among compounds in the same category, that is, \( \varepsilon > 1/2 \) virtually always in every organism, and therefore \( \gamma < 1 \). However, many mutation estimation studies focus only on the determination of \( q \) (see for instance [22], [23]). If only the parameter \( q \) is known, one may assume the simplification \( \gamma = 1 \). This is known as the Jukes-Cantor model in molecular evolution studies. In this model all off-diagonal entries of \( \Pi \) are equal, that is, \( p(Z = z | Y = y) = q/3 \) for all \( z \neq y \). Several observations are made for the capacity analysis with the Jukes-Cantor model throughout this paper.

The mutation model chosen implies a symmetric channel, since all rows (or columns) of \( \Pi \) contain the same four probabilities. Among all memoryless models used in molecular evolution, the Kimura model is the one with the highest number of parameters while still yielding a symmetric channel. It is well known that symmetric channels have a simple capacity analysis, which is exploited where possible. The most general time-reversible substitution mutations model may have up to 9 independent parameters, which in general yields a nonsymmetric channel. However, according to Li [24] mutation models with many parameters are not necessarily accurate, due to the estimation issues involved.

It will be assumed that the matrix \( \Pi \) models the base substitution mutations undergone by the genome of an organism during one generation. A Markov chain \( Y \to Z_{(1)} \to Z_{(2)} \to \cdots \to Z_{(m)} \) can be used to model \( m \) generations of an organism at a given site (position) in a genome, whose corresponding base at generation \( m = 0 \) is represented by \( Y \). The site can lie anywhere in the genome of an organism with asexual reproduction, or in the non-recombinant sections of an organism with sexual reproduction. The relationship between \( Y \) and \( Z_{(m)} \) is given by the transition probability matrix \( \Pi^m = [p(Z_{(m)} = z | Y = y)] \). As \( \Pi = \Pi^T \) one can write \( \Pi^m = V D^m V^T \), with the eigenvalues of \( \Pi \) arranged in a diagonal matrix \( D \triangleq \text{diag}(1, \lambda, \mu, \mu) \), where

\[
\lambda \triangleq 1 - \frac{4\gamma}{3}q, \quad (3)
\]

\[
\mu \triangleq 1 - 2(\frac{1}{3} - \sqrt{\frac{1}{3}})q, \quad (4)
\]

and \( V \) is a matrix whose columns are the normalised eigenvectors of \( \Pi \) associated with the corresponding eigenvalues in \( D \), that is

\[
V = \begin{bmatrix}
1 & 1 - \sqrt{\frac{1}{3}} & 0 & 0 \\
1 & 1 & 0 & -\sqrt{\frac{1}{3}} \\
1 & 1 & 0 & \sqrt{\frac{1}{3}} \\
1 & 1 & \sqrt{\frac{1}{3}} & 0
\end{bmatrix} \quad \text{(5)}
\]

From the diagonalisation of \( \Pi^m \) it follows that the elements of its diagonal take all the value \( \frac{1}{4}(1 + 2\mu^m + \lambda^m) \), the elements of its skew diagonal take all the value \( \frac{1}{4}(1 - 2\mu^m + \lambda^m) \), and the rest of its entries are \( \frac{1}{4}(1 - \lambda^m) \). Therefore any row (or column) of this matrix contains the same probabilities, as \( \Pi^m \) is also the transition matrix of a symmetric channel. From the diagonal elements one can see that the accumulated base substitution mutation rate after \( m \) generations is given by

\[
q^{(m)} = p(Z_{(m)} \neq y | Y = y) = 1 - \frac{1}{4}(1 + 2\mu^m + \lambda^m). \quad (6)
\]

It can be verified that \( \lim_{m \to \infty} q^{(m)}|_{\gamma>0} = 3/4 \), but \( \lim_{m \to \infty} q^{(m)}|_{\gamma=0} = 1/2 \). This is because \( |\mu| < 1 \) for any \( \gamma \) and \( |\lambda| < 1 \) when \( \gamma > 0 \), but \( \lambda = 1 \) when \( \gamma = 0 \). The behaviour of the particular case \( \gamma = 0 \) is connected to the fact that we must have both \( q \in (0,1) \) and \( \gamma \in (0,3/2) \) for the Markov chain to be aperiodic and irreducible, and thus possess a limiting stationary distribution. From the previous considerations, the limiting distribution —that is, the distribution of \( Z_{(\infty)} \)— is uniform, because \( \lim_{m \to \infty} \Pi^m = \frac{1}{4} I^T 1 \). When \( \gamma = 1 \) and \( q = 3/4 \) then \( \Pi = \frac{1}{4} I^T 1 \); hence \( Z_{(m)} \) is always uniformly distributed for any \( m \) as in the limiting case.
Lastly, under the base substitution mutation model considered, codons undergo a mutation channel modelled by the $64 \times 64$ transition probability matrix
\[
\Pi = [p(Z = z|Y = y)] = \Pi \otimes \Pi \otimes \Pi,
\]
where $\otimes$ is the Kronecker product. This is because $p(Z = z|Y = y) = \prod_{i=1}^m p(Z_i = z_i|Y_i = y_i)$ according to our memoryless channel assumption. Trivially this channel is also symmetric. When $m$ mutation stages are considered, $\Pi^m$ replaces $\Pi$ in (7), since $\Pi^m = [p(Z^{(m)} = z|Y = y)] = (\Pi \otimes \Pi \otimes \Pi)^m = \Pi^m \otimes \Pi^m \otimes \Pi^m$ [25].

III. CAPACITY ANALYSIS

A. Noncoding DNA

In this section we consider the embedding capacity of noncoding DNA, which also establishes a basic upper bound to capacity for protein-coding DNA. As discussed in Section II, we are assuming that the information embedding operation can either overwrite or append a host ncDNA strand $\chi$, which amounts to freely choosing the input $y$ to the mutation channel. Therefore with ncDNA the embedding capacity is simply given by $C_{nc} \triangleq \max I(Z^{(m)}; Y)$ bits/base, where the maximisation is over all distributions of $Y$. This capacity is that of a symmetric channel for the mutation model considered. With this type of channel $H(Z^{(m)}|Y)$ is independent of the input, and uniformly distributed $Y$ leads to uniformly distributed $Z^{(m)}$. Hence
\[
C_{nc} = \log |\chi| - H(Z^{(m)}|Y) \text{ bits/base},
\]
where $H(Z^{(m)}|Y)$ is the entropy of any row of $\Pi^m$, that is,
\[
H(Z^{(m)}|Y) = -\sum_{z \in \chi} p(Z^{(m)} = z|Y = y) \log p(Z^{(m)} = z|Y = y)
\]
for any $y \in Y$. $C_{nc}$ is nonincreasing on $m$, since the information processing inequality implied by the Markov chain $Y \rightarrow Z_1 \rightarrow Z_2 \rightarrow \cdots \rightarrow Z^{(m)}$ implies that $I(Z_1; Y) \geq I(Z_2; Y) \geq \cdots \geq I(Z^{(m)}; Y)$. As long as the Markov chain is aperiodic and irreducible then $\lim_{m \rightarrow \infty} C_{nc} = 0$. The reason for this behaviour is that the limiting distribution is in this case independent of $Y$, which implies that $\lim_{m \rightarrow \infty} H(Z^{(m)}|Y) = H(Z^{(\infty)}) = \log |\chi|$. It is interesting to note that, under aperiodicity and irreducibility of the Markov chain, this zero limiting capacity will also apply to models more involved than (1), such as those in which the channel matrix is parameterised by up to 9 independent values. Lastly, we also have $C_{nc}|_{\gamma=1,q=3/4} = 0$, since with these parameters $Z^{(m)}$ is always uniformly distributed.

As a function of $\gamma$ the ncDNA capacity is bounded for every $m$ and $q$ as follows
\[
C_{nc}|_{\gamma=1} \leq C_{nc} \leq C_{nc}|_{\gamma=0}.
\]
Although it can be shown that these inequalities hold true in general, it is much simpler to prove for the range of interest $\gamma \leq 1$ and $q \leq 1/2$. The latter condition implies that both $0 \leq \lambda \leq 1$ and $0 \leq \mu \leq 1$. For fixed $m$ and $q$, the maximum (respectively, minimum) of $C_{nc}$ over $\gamma$ corresponds to the minimum (respectively, maximum) of the accumulated base mutation rate $q^{(m)}$. Differentiating (6) we obtain $\partial q^{(m)}/\partial \gamma = (mq/3) (\lambda^{m-1} - \mu^{m-1})$. Therefore $q^{(m)}$ is monotonically increasing when $\gamma \leq 1$ (as this corresponds to $\lambda \geq \mu$), and then its maximum in that range occurs when $\gamma = 1$ and its minimum when $\gamma = 0$.

The upper bound can be written as
\[
C_{nc}|_{\gamma=0} = 2 - h \left( \frac{1}{2} + \frac{1}{2} (1 - 2q)^{m} \right).
\]
Notice that $\lim_{m \rightarrow \infty} C_{nc}|_{\gamma=0} = 1$, that is, the capacity limit is not zero when $\gamma = 0$ because then the Markov chain is reducible. $\gamma = 0$ cannot occur in reality since it would imply that transversion mutations are impossible. However it illustrates that the higher the transition/transversion ratio $\varepsilon$, the higher the capacity.

Figures 1 and 2 show $C_{nc}$ for two different values of the base substitution mutation rate per replication, $q = 10^{-5}$ and $q = 10^{-9}$. These mutation rates have been arbitrarily chosen to illustrate the capacity values in representative cases of $q$ corresponding to viruses [23] and microbes [22, Table 4], respectively. The results in the figures show the validity of the bounds (9) and the limiting behaviours discussed. From these figures one can also empirically see that, as a rule-of-thumb, capacity lies around $C_{nc} \approx 10^{-2}$ bits/base for $m \approx 6/(5q)$. Finally, in Figure 3 we compare the ncDNA capacity computed with the Kimura model versus the same computation using an unconstrained estimate of $\Pi^m$ given by Li [24,
are uniformly distributed, using the average \( C_{nc} = 10^{−4} \text{bits/base} \). The basic effect of accumulated mutations on capacity (exponential decrease with mutation rate) was numerically evaluated by May et al. [28], using values of \( q \) estimated from different organisms and the Jukes-Cantor and Kimura \((\gamma = 1/2)\) models. Some authors have also considered the behaviour of capacity for \( m > 1 \). Gutfraind [29] discussed the basic effect of accumulated mutations on capacity (exponential decrease with \( m \)), but used a binary alphabet and the binary symmetric channel. Both Battail [18] and May [17] computed capacity under cascaded mutation stages using a quaternary alphabet and the Jukes-Cantor model. Battail obtained his results analytically —using a continuous-time approach rather than the discrete-time approach taken in this paper— and May obtained hers numerically. The results by Battail are essentially consistent with the ones presented here, but the results by May are not. However it would appear that this effect is due to the use of \( n' = m \times g \) as the number of generations in [17], where \( g \approx 3 \times 10^9 \) base pairs is the length of the human genome.

In any case, none of the aforementioned approaches reflect the capacity increase afforded by a mutation model allowing \( \gamma < 1 \). It is possible that the trend towards higher capacity observed as \( \gamma \to 0 \) implies that evolution has favoured genetic building blocks which feature an asymmetric behaviour under mutations (that is, in the current genetic machinery pyrimidines versus purines is not the case of a hypothetical perfectly mutation-symmetric set of four bases for which \( \gamma = 1 \)). If this assumption is correct, this information-theoretical induced “mutation-symmetry breaking” must have occurred early in evolutionary terms, since it is widely believed that the current genetic machinery evolved from a former “RNA world” [30] in which life would only have been based on the self-replicating and catalysing properties of RNA. In the RNA world there would not have been translation to proteins, and therefore no genetic code, and hence information was freely encoded using a 4-ary alphabet almost exactly like the one used in DNA. Note that uracil, which replaces thymine in RNA, is also a pyrimidine, that is, in the RNA world \( Y = \{ U, C \} \).

With these facts in mind, we may model the maximum transmissible information under mutations in the RNA world by relying on (8), and thus see that the mutation-symmetry breaking conjecture above applies to the evolution of RNA from hypothetical ancestor genetic building blocks. One must bear in mind that single-stranded molecules, such as RNA, are much more mutation-prone than double-stranded ones such as DNA. Therefore smaller values of \( m \) would have sufficed for some type of mutation-symmetry breaking to be relevant in terms of information transmission at early stages of life.

B. Protein-Coding DNA

Unlike in the ncDNA case, embedding information in pcDNA is a coding problem with side information at the encoder. This side information is the exact knowledge by the encoder of the primary structure of the protein encoded by the host \( X \), that is, \( X' = \alpha(X) \). The primary structure \( X' \) determines the encoder state because it must hold for the information-carrying sequence \( Y \) that \( \alpha(\gamma_i) = \alpha(x_i) = x'_i \) for all \( i \). This is equivalent to hiding data in a discrete host under an embedding constraint. Nevertheless, apart from the trivial difference of using a 4-ary instead of a typical 2-ary alphabet, several issues distinguish pcDNA data embedding as a special problem.

In order to illustrate these issues consider momentarily a typical data hiding scenario in which a discrete binary host, that is \( X = [x_1, \cdots, x_n] \) with \( x_i \in \mathcal{X} = \{0, 1\} \), is modified to embed a message \( b \) from a given alphabet. The watermarked signal \( Y = e(X, b) \) must be close to \( X \), where closeness is usually measured by means of the Hamming distance \( d_H(Y, X) \). Pradhan et al. [31] and Barron et al. [32] have determined the achievable rate in this scenario, assuming that the elements of \( X \) are uniformly distributed, using the average distortion constraint \( \frac{1}{n} E[d_H(Y, X)] \leq d \), and supposing that \( Y \) undergoes a memoryless binary symmetric channel with crossover probability \( q \). Their result is

\[
R^\text{unif} = \text{u.c.e.}\{h(d) - h(q)\} \text{ bits/host symbol},
\]

where \( \text{u.c.e.}\{\cdot\} \) is the upper concave envelope of the argument (when constrained to nonnegative values). Similarly, our initial

\[\text{For instance, the genomes of RNA viruses such as HIV are known to exhibit as a whole base substitution mutation rates up to } q = 8 \times 10^{-5} \text{ [23], whereas the genomes of DNA-based organisms typically exhibit as a whole base substitution mutation rates } 10^{-3} \text{ times lower.}\]
goal for pcDNA data embedding is to obtain the achievable rate $R_{pc}^{X'}$ for a fixed distribution of $X'$ under the symmetric channel discussed in Section II-A, in particular when $X$ is uniformly distributed, which is denoted as $R_{pc}^{(uni)}$, similar to the analyses of Pradhan et al. [31] and Barron et al. [32]. Furthermore we will also determine capacity, namely, the maximum achievable rate over all distributions $X'$ of the primary structure encoded by the host.

The first important difference in the pcDNA data embedding scenario is that average inequality constraints on the Hamming distance—such as the ones used in [31], [32]—are meaningless if one wants to carry through to $\tilde{Y}$ the full biological functionality of $X$. Instead, since it must always hold that $\alpha(\tilde{Y}) = \alpha(X)$, one must establish the deterministic constraint

$$d_H(y', x') = \sum_{i=1}^{n} d_H(y'_i, x'_i) = 0. \quad (11)$$

This is equivalent to assuming that the primary structures of the proteins encoded by the information-carrying sequence and by the host sequence must be identical. This requires that $d_H(y'_i, x'_i) = 0$ for all $i = 1, \ldots, n$.

The second distinguishing feature of pcDNA data embedding is due to the variable support of the channel input variable. Whereas in discrete data hiding with binary host, one always has $y_i \in \{0, 1\}$ independently of $x_i$, in pcDNA data embedding we must have $y_i \in S_{u}(x_i)$ so that the constraint (11) can always be satisfied. Therefore the support of $y_i$ is dependent on $x_i$, as codon equivalence is not evenly spread over the ensemble of amino acids (see Table I).

1) Achievable Rate: Since side information at the encoder must be taken into account in the pcDNA case, then the achievable rate is given by Gel’fand and Pinsker’s formula [33] $R_{pc}^{X'} = \max \{ I(Z(m)|U) - I(X'; U) \}$ bits/codon, where $X'$ is the distribution of the primary structure that acts at the side information at the encoder, $U$ an auxiliary random variable which must be determined for the problem at hand, and the maximisation is for nonnegative values of the functional on all distributions $p(y, u|x')$ under the biological constraint $d_H(\alpha(y), x') = 0$. The intuition behind $U$ is that this variable should simultaneously act as a good source code for representing the side information (by minimising $I(X'; U)$ within the distortion constraint), and as a good channel code for conveying information over the mutation channel (by maximising $I(Z(m)|U)$).

Gel’fand and Pinsker showed in [33] that in the maximisation problem above one may assume that the channel input $Y$ is a deterministic function of the side information $X'$ and the auxiliary variable $U$, that is, $\tilde{Y} = e(X', U)$. Consequently the relationship between the variables in the problem can be summarised by the diagram in Figure 4. Since the support of $Y|x'$ can only be the set of codons $S_{x'}$ —so that the biological constraint $d_H(\alpha(y), x') = 0$ can always be met—then the cardinality of the support set of $U|x'$ must coincide with the multiplicity of $x'$, that is, $|S_{x'}|$. As $U$ is an arbitrary auxiliary variable, we may choose the support set of $U|x'$ to be $S_{x'}$. This allows us to define the embedding function $Y|x' = e(x', U)$ as $Y|x' \equiv U|x'$ without loss of generality. Any of the $|S_{x'}|!$ permutations of the elements of $S_{x'}$ could actually be chosen to establish this function.

From these considerations it follows that $Y = U$. As a result it is assumed in the subsequent computations that $U$ becomes the mutation channel input (equivalently, the encoder output). Noticing that $S_{x'} \cap S_{y'} = \emptyset$ for $x' \neq y'$, the distribution of $U$ can be put as $p(u) = p(u|x')p(x')$ when $u \in S_{x'}$, because $p(v|x') = 0$ when $v \notin S_{x'}$. This discussion on $U$ also implies that $H(X'|U) = 0$, since given a codon $u$ there is no uncertainty on the amino acid represented by it, and therefore $I(X'; U) = H(X')$.

Since $Y|x', u$ is deterministic, the maximisation on $p(y, u|x')$ amounts to a maximisation on $p(u|x')$. Hence the achievable rate for a given distribution of $X'$ can be expressed as

$$R_{pc}^{X'} = \max_{p(u|x')} I(Z(m)|U) - H(X') \text{ bits/codon.} \quad (12)$$

As $H(Z(m)|U)$ only depends on the transition probabilities of the symmetric channel given by $\Pi^n$, and as trivially $H(X')$ only depends on $X'$, the optimisation in (12) amounts to the constrained maximisation of $H(Z(m))$ on $p(u|x')$.

For the same reasons as the bounds (9) in Section III-A, the rate (12) is bounded as $R_{pc}^{X'}|\gamma=1 \leq R_{pc}^{X'} \leq R_{pc}^{X'}|\gamma=0$. Also, there are several cases where $R_{pc}^{X'}$ can be analytically determined, which are discussed next. First of all, since $C_{nc}|\gamma=1, q=3/4 = 0$ then $R_{pc}^{X'}|\gamma=1, q=3/4 = 0$ for any $X'$, because it must hold that $R_{pc}^{X'} \leq 3C_{nc}$. Therefore in this catastrophic case the choice of $p(u|x')$ is irrelevant. Furthermore it can be shown that $p(u|x') = 1/|S_{x'}|$ —that is, $U|x'$ uniformly distributed—is the exact maximising strategy in two situations, which are discussed in the following two lemmas.

**Lemma 1.** If $q = 0$ then the achievable rate is

$$R_{pc}^{X'}|q=0 = E[\log |S_{X'}|] \text{ bits/codon.} \quad (13)$$

**Proof:** Using the chain rule of the entropy one can write $H(U, X') = H(U) + H(X'|U) = H(X') + H(U|x')$. Since $H(X'|U) = 0$, and as $Z_{(m)} = U$ when $q = 0$, then the achievable rate is given by $R_{pc}^{X'}|q=0 = \max_{p(u|x')} H(U) - H(X') = \max_{p(u|x')} H(U|x')$. As $H(U|x') = \sum_{x' \in X'} p(x') H(U|x')$ is maximised when $H(U|x')$ is maximum for all $x'$, then $U|x'$ must be uniformly distributed in all cases. Then $H(U|x') = \log |S_{x'}|$ and (13) follows.
Remark. Note that (13) is the embedding rate intuitively expected in the mutation-free case. For example, if $X$ were uniformly distributed, which would yield $X' = \alpha(X)$ nonuniform with pmf $p(x') = |S_{x'}|/|X|^3$, then one would intuitively compute the rate as $R^{\text{pc}(\text{unif})}_{\text{nc}}|_{q=0} = \sum_{x'} |S_{x'}| \log |S_{x'}| = 1.7819 \text{ bits/codon}$, since $|S_{x'}|$ choices are available to the embedder when the host amino acid is $x'$. The rate when $X$ is uniform can actually be obtained in closed form for every $q$ using the following result.

**Lemma 2.** If $X$ is uniformly distributed then the achievable rate is

$$R^{\text{nc}(\text{unif})}_{\text{pc}} = \bar{C}_{\text{nc}} - H(X') \text{ bits/codon},$$

where $\bar{C}_{\text{nc}} = \max I(Z_{(m)}; U)$ and this maximisation is unconstrained on $p(u)$, that is, $\bar{C}_{\text{nc}}$ is the capacity of the symmetric codon mutation channel with transition probability matrix $\Pi^m$.

**Proof:** Since $p(u) = p(u|x')p(x')$ when $u \in S_{x'}$, if $X$ is uniformly distributed then $p(u) = p(u|x')|S_{x'}|/|X|^3$ when $u \in S_{x'}$. Therefore choosing $U|x'$ to be uniformly distributed implies that $p(u) = 1/|X|^3$ for all $u$. Since $\Pi^m$ represents a symmetric channel and a uniform input maximises mutual information over such a channel, then $\bar{C}_{\text{nc}} = \max p(u) I(Z_{(m)}; U)$ is achieved in (12).

**Remarks.** Since $\bar{C}_{\text{nc}|q=0} = \log |X|^3$, it can be seen that the particular case discussed in the previous remark can also be written as $R^{\text{pc}(\text{unif})}_{\text{nc}}|_{q=0} = \log |X|^3 - H(X')$. An interesting insight is provided by the fact that the three parallel symmetric channels undergone by the three bases in a codon are mutually independent, and as a result one can use the equality $\bar{C}_{\text{nc}} = 3C_{\text{nc}}$ in (14). As $H(X')$ is the lower bound to the lossless source coding rate of $X'$, expression (14) tells us something which is intuitively appealing but only exact when $X$ is uniform: the pcDNA embedding rate equals three times the ncDNA embedding rate minus the rate needed to losslessly convey the primary structure encoded by the host to the decoder.

Unlike the case above, the distribution of $X$ in real pcDNA sequences (that is, genes) is not uniform. To start with, there can only be a single $Stp$ codon in a sequence that encodes a protein, which rules out uniformity of $X$. For this reason we show next how to compute the achievable rate in the general case, which corresponds to an arbitrary host sequence encoding a primary structure distributed as $X'$. As in many other channel capacity problems, it does not seem possible to analytically derive a general optimum set of conditional pmf’s $p(u|x')$ in order to compute the maximum achievable rate (12). One can pose the analytical optimisation problem and see that it involves solving a nontrivial system of $|X|^3 + |X'|$ nonlinear equations and unknowns. However the numerical solution is straightforward by means of the Blahut-Arimoto algorithm [26] adapted to a side-informed setting. Such an algorithm has been described by Dupuis et al. [34] to determine the rate given by the general Gel’fand and Pinsker formula. However the Blahut-Arimoto algorithm can be given

\[^4\alpha(\text{unif}) \text{ denotes the distribution of } X' \text{ when } X \text{ is uniformly distributed.}\]
share their two first bases, if \( q \ll 1 \) we can produce the approximations
\[ p(z|x') \approx 0 \text{ when } z \notin S_{x'}, \]
and
\[ p(z|x') \approx \left( \sum_{v \in S_{x'}} p(z|v) \right) \left( \sum_{u \in S_{x'}} p(z|u) p(u|x') \right) \text{ when } z \in S_{x'}.
\]
With this simplification, whenever \( \sum_{u \in S_{x'}} p(z|u) \) is constant for all \( z \in S_{x'} \), choosing \( p(u|x') \) to be uniform implies that \( p(z|x') \) is also uniform, which maximises \( H(Z) \text{| } x' \).

It can be verified that this condition holds for all \( x' \) such that \( |S_{x'}| = 1, 2, 4 \), which accounts for 16 out of the 21 elements in \( X' \).

Figures 6–9 present the achievable rates in pcDNA for several distributions of \( X' \); for the same two values of \( q \) in Section III-A and for \( \gamma = 1 \) (Jukes-Cantor model) and \( \gamma = 0.1 \) (realistic case). Two of the \( X' \) distributions were empirically obtained from the amino acid sequences encoded by two real genes: Ypt7 (\( S. \ Cerevisiae \)) and FtsZ (\( B. \ Subtilis \)), whose GenBank accession numbers are NC_001145 and NC_000964, respectively. Also depicted are the rate (14) for \( X \) uniform and the rate for the deterministic distribution of \( X' \) with outcome Ser, which, as we will discuss in Section III-B2, yields capacity. The plots show that the difference between the results obtained with the Blahut-Arimoto algorithm and those obtained with \( p(u|x') = 1/|S_{x'}| \) is negligible. In any case, all achievable rates are well below \( 3C_{nc} \) bits/codon, the rate that can be attained when the embedder is unconstrained. The most important observed effect is that the rates for the two real genes are not far away from the rate for the uniform case, which has the advantage of being analytically determined. There is also a clear cutoff point beyond which no information can be reliably embedded within real genes, which, as an empirical rule-of-thumb, lies around \( m \approx 1/(10q) \). Lastly the capacity results are higher for the realistic case with respect to the Jukes-Cantor model, similar to the results in Section III.

a) Codon bias preservation: A final consideration is that real genes exhibit codon bias, that is, distributions of \( X | x' \) characteristic to particular species. However the maximisation in (12) implies that, in general, the codon bias of the host sequence \( X \) will not be maintained in the optimum information-carrying sequence \( Y \). Although both sequences will still translate into the very same protein (since \( \alpha(Y) = \alpha(X) \text{ is enforced} \), if their codon biases are very different
this can have a marked influence on biologically relevant features such as gene translation times [19]. This subsection analyses the achievable rate when an additional codon bias preservation (cbp) constraint is enforced, as this can enhance the biocompatibility of pcDNA data embedding.

The basic observation to be made is the following: if one requires that the original codon bias of the host sequence be preserved in the information-carrying sequence, then one must peg the distribution of $U|\gamma(x'|x')$ to the corresponding conditional distribution of the host sequence. With this special constraint, no maximisation on $p(u|x')$ is required, or possible, and as $p(U = u|X = u)$ is the corresponding rate is just $R_{pc, cbp} \triangleq I(\mathcal{Z}^{(m)}; X') - H(X') \leq R_{pc}'$. The codon bias preservation constraint is equivalent to a steganographic constraint in data hiding problems, since the pmf of the host codons is preserved in the information-carrying sequence. This mirrors Cachin’s criterion for perfect steganography [36]. A comparison of maximum rates and codon bias preservation rates for the same two real genes is given in Figure 10. Note that $R_{pc, cbp} = R_{pc}(umf)$ because of Lemma 2.

In reality preserving the codon bias of an actual host sequence implies preserving its codon count, that is, its codon histogram. Genes are usually short, typically in the order of hundreds of codons, and thus the degrees of freedom for preserving the codon count in the information-carrying gene may not be high. Consequently the information-theoretical analysis for codon bias preservation will usually be optimistic regarding the actual achievable rate when enforcing codon count preservation (ccp).

It is therefore interesting to analyse what can actually be achieved in the latter case. It is assumed in the following that $q = 0$, that is to say, a mutation-free scenario. Let $n_\sigma$ be the number of times that codon $\sigma$ appears in an $n$-codon long gene $\mathbf{X} = [x_1, x_2 \cdots, x_n]$. Therefore $n = \sum_{\sigma \in \Sigma} n_\sigma$. The number $r$ of different DNA sequences with the same amino acid translation and the same codon count as $\mathbf{X}$ is given by the following product of multinomial coefficients:

$$r = \frac{(\sum_{\sigma \in \Sigma} n_\sigma)!}{\prod_{\sigma \in \Sigma} n_\sigma!}.$$  \hspace{1cm} (15)

This is because for each amino acid $\sigma$ the multinomial coefficient gives all possible rearrangements of codons corresponding to $\sigma$ at its $n_\sigma$ positions in the sequence, which keeps the codon count fixed. Therefore the mutation-free codon count preservation rate for the finite sequence $\mathbf{X}$ is $R_{pc, cbp}^{\text{cbp}}|q=0 = (1/n) \log r$ bits/codon. If $\mathbf{X}$ represents the empirical distribution of codons in $\mathbf{X}$ then one clearly has $R_{pc, cbp}^{\text{cbp}}|q=0 \leq R_{pc, cbp}^{\text{x}}|q=0$, and the rate on the right of the inequality can only be achieved asymptotically, as the length of the sequence distributed as $\mathbf{X}$ goes to infinity. Interestingly both rates are close for real genes. For instance, genes FtsZ and Ypt7 can be deemed to be short: 386 and 208 codons long, respectively. Despite this, $R_{pc, cbp}^{\text{cbp}}|q=0 = 1.3723$ bits/codon and $R_{pc, cbp}^{\text{x}}|q=0 = 1.2797$ bits/codon, which are just slightly lower than the asymptotic rates for the same genes, $R_{pc, cbp}^{\text{cbp}}|q=0 = 1.4201$ bits/codon and $R_{pc, cbp}^{\text{x}}|q=0 = 1.3179$ bits/codon.

In any case, it is shown in Appendix B that for large $n$

$$R_{pc, cbp}^{\text{cbp}}|q=0 \approx R_{pc, cbp}^{\text{x}}|q=0 = H(\mathbf{X}) - H(X') \text{ bits/codon},$$

simply put, the combinatorial rate for codon count preservation coincides asymptotically with the Gel’fand and Pinsker’s formula with codon bias preservation.

2) Capacity: To conclude, we undertake the computation of the capacity of pcDNA data embedding, which is the maximum achievable rate over all possible distributions of the primary structure encoded by a gene. This problem can be put as

$$C_{pc}^* \triangleq \max_\mathbf{p} R_{pc}^{X'} \text{ bits/codon.}$$

The explicit computation of $C_{pc}^*$ is simple in two particular cases considered in Section III-B1 already:

- $\gamma = 1$ and $q = 3/4$: Trivially, $C_{pc}^*|\gamma=1,q=3/4 = R_{pc}^{X'}|\gamma=1,q=3/4 = 0$. However, the point that needs to be made is that although this is true for any $X'$, only deterministic $X'$ yields $H(X') = 0$ exactly, and so, by the continuity of the rate functional, this is the best strategy when approaching $q = 3/4$ from the left.

- $q = 0$: From Lemma 1, $C_{pc}^*|q=0 = \max_{y'} |S_{y'}|$ since $|S_{y'}| = 6$ is the maximum for all $x' \in \mathcal{W}' \triangleq \{\text{Ser, Leu, Arg}\}$, a distribution of $X'$ that maximises (13) is any for which $\sum_{x' \in \mathcal{W}'} p(x') = 1$. Note that $X'$ may be deterministic, but it does not have to be so. Capacity is then $C_{pc}^*|q=0 = \log 6 = 2.5850$ bits/codon.

**Remark.** A trivial upper bound for any $q$ is $C_{pc} \leq C_{pc}^*|q=0$. Since $C_{pc}^*|q=0 < 3 C_{pc}^*|q=0 = 6$, then side-informed pcDNA data embedding capacity will not be able to achieve non-side-informed ncDNA capacity for every mutation rate. This parallels the results in side-informed encoding with discrete hosts by Pradhan et al. [31] and Barron et al. [32], which have already been discussed. However it does not parallel the well-known “writing on dirty paper” result by Costa [37], which corresponds to continuous Gaussian host, mean squared error distortion, and additive independent Gaussian channel.
From our discussion above on $C_{pc}$ for $q = 3/4$ (with $\gamma = 1$) and for $q = 0$, one might surmise that a primary structure pmf with support in $\mathcal{W}'$ could be capacity-achieving in all cases. The actual capacity-achieving distribution is given by the following theorem:

**Theorem 1.** Capacity is achieved by the deterministic pmf of $X'$ that maximises $H(\mathbf{Z}(m))$.

**Proof:** See Appendix C.

**Remarks.** Denoting as $\xi'$ the deterministic outcome of $X'$, it can be numerically verified that $\xi' = \text{Ser}$ maximises $H(\mathbf{Z}(m))$ for all $\gamma \leq 1$, $m$, and $q$, and thus $C_{pc} = R^{\text{Ser}}_{pc}$ in these conditions. Some examples showing this fact and the rates achievable with deterministic $X'$ are presented in Figures 11–14, which are obtained for the same $\gamma$ and $q$ parameters as Figures 6–9. The maximum rates, corresponding to Ser in all cases, are highlighted. These plots also show that the rates obtained using the linearised approximation given in Appendix C are practically indistinguishable from the rates obtained using the Blahut-Arimoto algorithm. The approximation $p(u|x') = 1/|S_z|$ is also good, but it worsens as $\gamma$ decreases.

a) **Biological interpretations:** The results in this section are only concerned with artificial information embedding in pcDNA, and thus appear to have less obvious consequences in biological terms than the results concerning ncDNA. However an intriguing phenomenon which may allude to a biological meaning of these achievable rates is observed in Figures 13 and 14, that is, the range of $\gamma$ in which the model is realistic. This phenomenon consists of a rate drop for two particular values of $\xi' \in \mathcal{X}'$ as $m \to \infty$ with respect to all other symbols with the same multiplicity $|S_{\xi'}|$. These two values of $\xi'$ correspond to the stop symbol (Stp) and to the amino acid Leu, two codons of which happen to have the twin function of start codons in eukaryotic organisms. Therefore, when considered as individual information transmission blocks, all stop codons and most of the start codons seem to be less suited than the rest of codons for carrying extra information (redundancy). One may surmise that these codons may have undergone special evolutionary pressures during the emergence of the genetic code, due to their very specific functions. One possibility may be that since the stop symbol cannot be repeated in a single gene, it has been under less pressure to perform as a genetic information carrier. A similar effect may have been at work in the evolution of the start codons corresponding to Leu. However a more interesting conjecture is that the narrower “genetic windows” for these special codons are due to natural selection broadly favouring an increase of the complexity of organisms over the generations, which by and large requires gene lengthening. Making these special codons less reliable, plus the effect of natural selection, may have been a mechanism for such progressive elongation of genes.

**IV. Conclusion**

This paper has provided an analysis of the Shannon capacity of DNA data embedding when mutations are described according to the Kimura model of molecular evolution. The analysis may be used to assess the optimality of a growing number of bioengineering procedures aimed at inserting arbitrary information in the genomes of living organisms, both within noncoding and protein-coding sections.

Some biological connections of the results given have also been discussed, leading to some unexpected insights into the mutation-asymmetry of the DNA bases and into the behaviour of key building blocks in the genetic code —start and stop codons. These results indicate that the Shannon capacity may hold clues to a deeper understanding of the genetic machinery, which is an exciting possibility.

Further research should consider insertion and deletion mutations (indels) in order to complete the treatment of the errors that a DNA data embedding algorithm may face. Although the exact capacity computation under indels is an unsolved problem in most digital communications settings, bounding strategies and approximations reliant on realignment methods which are widely used in bioinformatics may be useful for addressing this problem. Generalisations of the Kimura model with up to nine parameters may also be considered. However it has been shown by means of an unconstrained real estimate of the mutation channel that generalised models may not overly change the analysis herein.

**APPENDIX**

A. **Blahut-Arimoto algorithm for pcDNA data embedding**

Since the subtracting entropy in (12) is independent of $p(u|x')$ we just need to maximise $I(\mathbf{Z}(m), \mathbf{U})$. In the standard Blahut-Arimoto algorithm one would alternatively maximise $p^\ast(u|z)$ and $p^\ast(u)$ using $p(z|u)$ (that is, $\mathbf{U}^m$). However the distribution of $\mathbf{U}$ is constrained in (12) by the distribution $p(x')$ of the side information, and instead we have to maximise $p^\ast(u|z)$ and $p^\ast(u|x')$ alternatively. Writing the mutual information as

$$I(\mathbf{Z}, \mathbf{U}) = \sum_{x' \in \mathcal{X}'} \sum_{z \in \mathcal{X}^3} p(z|u)p(u|x')p(x') \log \frac{p(u|z)}{p(u|x')p(x')}$$

(18)

it is straightforward to see (following the same reasoning as in [26]) that the two iterative maximisation steps in this case are

1. **Step 1:**

   $$p^\ast(u|z) = \frac{p(z|u)p(u|x')p(x')}{\sum_{y'} \sum_{v \in S_{y'}} p(z|v)p(v|y')p(y')}$$

   (19)

   where $u \in S_{x'}$.

2. **Step 2:**

   $$p^\ast(u|x') = \frac{\prod_{z} p(u|z)p(z|u)}{\sum_{v \in S_{y'}} \prod_{z} p(v|z)p(z|v)}.$$  

(20)

Note that the first step is the same as in the standard Blahut-Arimoto algorithm, because $p(u) = p(u|x')p(x')$ if $u \in S_{x'}$. The second step is also essentially the same, but it is applied $|\mathcal{X}'|$ times to determine each input distribution $p(u|x')$ instead of only once to determine the input distribution $p(u)$. 


Assuming that \( n \) is large, this amount may be developed using Stirling’s approximation \( \log t! \approx t \log t - t \) (for natural logarithms). When using this approximation in (21) the summands without a logarithm cancel out and we get:

\[
l \approx \sum_{x'} \left( \sum_{x \in S_{x'}} n_x \right) \log \left( \sum_{x \in S_{x'}} n_x \right) - \sum_{x'} \sum_{x \in S_{x'}} n_x \log n_x.
\]

(22)

Using the asymptotic expressions \( p(x') \approx (1/n) \sum_{x \in S_{x'}} n_x \) and \( p(x) \approx (1/n) n_x \), it can be seen after some algebra that (22) can be in turn approximated by

\[
l \approx n \left( \sum_{x'} p(x') \log p(x') - \sum_{x' \in X'} \sum_{x \in S_{x'}} p(x) \log p(x) \right).
\]

(23)

Therefore the rate \( R_{pc}^{\infty} \big|_{\gamma=0} = l/n \) can be asymptotically approximated for large \( n \) as \( R_{pc}^{\infty} \big|_{\gamma=0} \approx H(X) - H(X') \) bits/codon.

### C. Capacity-achieving strategy \( p(x') \)

In order to find the capacity-achieving strategy for the amino acid distribution one needs to solve

\[
\frac{\partial}{\partial p(x')} \left[ H(Z_{(m)}) - H(X') + \nu \left( \sum_{y' \in X'} p(y') - 1 \right) \right] = 0
\]

(24)
for \( x' \in X' \), with \( \nu \) a Lagrange multiplier. In what follows, \( p(z|x') \) denotes \( p(Z_{(m)} = z | X' = x' ) \) for notational convenience. For simplicity, and without loss of generality, we assume natural logarithms in the optimisation. Using \( \frac{\partial p(z)}{\partial p(x')} = p(z|x') \), expression (24) becomes

\[
\sum_{z \in X^3} p(z|x') \log \left( \sum_{y' \in X'} p(y') p(z|y') \right) = \log p(x') + \nu, \tag{25}
\]

for \( x' \in X' \). The solution remains unchanged if we multiply (25) across by \( p(x') \). This operation allows us to see by inspection that any extreme of the Lagrangian inside the differential in (24) has to be deterministic, that is, \( p(x') = 1 \) for some \( x' = \xi' \) and \( p(x') = 0 \) for \( x' \neq \xi' \). Note that this is in agreement with the strategies for the cases \( q = 0 \) and \( \gamma = 1 \) with \( q = 3/4 \) discussed in Section III-B2. One may verify that a uniform distribution of \( X' \) cannot possibly solve (25) for all \( x' \in X' \), because \( \sum_{x} p(z|x') \log \left( \sum_{y'} p(z|y') \right) \) is not constant on \( x' \) unless \( \gamma = 1 \) and \( q = 3/4 \), in which case we have shown that the Shannon capacity is zero for any distribution.

According to the previous discussion, for any capacity-achieving solution it always holds that \( H(X') = 0 \). Therefore we just have to maximise \( H(Z_{(m)}) \) over the ensemble of 21 deterministic distributions of \( X' \). The computation of \( R_{pe}^{\xi} \) and the maximising distribution \( U_{\xi'} \) can be accomplished using the Blahut-Arimoto algorithm, following the discussion in Section III-B1 on the optimal strategy for fixed \( p(x') \). Note that \( \xi' = \text{Trp} \) and \( \xi' = \text{Met} \) can be completely ruled out, since \( |\text{Trp}| = |\text{Met}| = 1 \), and then only null rates are possible in these cases. We only need therefore to compute \( R_{pe}^{\xi'} \) for the remaining 19 amino acids and then choose the maximum. Also, \( \xi' = \text{Stp} \) can be alternatively hypothesised, since this symbol can only appear exactly once in a gene.

\( \text{a) Approximation to maximising strategy for deterministic } X' \): It is also possible to provide a closed-form approximation to the maximising distribution \( U_{\xi'} \), which is more accurate than just using the approximation \( p(\text{u} | \xi') = 1/|S_{\xi'}| \) discussed in Section III-B1. It must firstly be observed that if \( X' \) is deterministic, then the side-informed setting amounts to a non-side informed setting with \(|S_{\xi'}|\) inputs and \( |\chi|^{3} \) outputs. This setting can be modelled by a transition probability matrix \( \Lambda \) whose rows are the rows of \( \Pi^{m} \) corresponding to the codons associated with \( \xi' \). In general this channel will not be symmetric nor weakly symmetric, since although its rows are permutations of the same set of probabilities, its columns are not, and their sum is not constant either. However \( H(Z_{(m)}|U) \) is still independent of the distribution of \( U \), so we only need to maximise \( H(Z_{(m)}) \) to find capacity. The corresponding conditions for the maximum are

\[
\sum_{z \in X^3} p(z|\text{v}) \log p(z) + 1 = \rho, \tag{26}
\]

for \( \text{v} \in S_{\xi} \), and with \( \rho \) a Lagrange multiplier. Using \( \log x \leq x - 1 \) and \( p(z) = \sum_{u \in S_{\xi}} p(z|u) p(u|\xi') \), one can write

\[
\sum_{z \in X^3} p(z|\text{v}) \sum_{u \in S_{\xi'}} p(z|u) p(u|\xi') \leq \rho, \tag{27}
\]

for \( \text{v} \in S_{\xi'} \). Our approximation consists of solving \( p(\text{u}|\xi') \) by enforcing equality in (27) for all \( \text{v} \in S_{\xi'} \). This yields the linear system

\[
\pi \left( \Lambda \Lambda^{T} \right) = \rho 1, \tag{28}
\]

where the probabilities \( p(\text{u}|\xi'), \) with \( \text{u} \in S_{\xi'} \), are the elements of the \( 1 \times |S_{\xi'}| \) vector \( \pi \) (arranged in the same codon order as the rows of \( \Lambda \)), and \( 1 \) is an all-ones vector of size \( 1 \times |S_{\xi'}| \). Since \( \pi \) must be a pmf, we may fix any arbitrary value of \( \rho \), such as \( \rho = 1 \), and then normalise the solution \( \tilde{\pi} \) to the resulting linear system, that is

\[
\tilde{\pi} = 1(\Lambda \Lambda^{T})^{-1}. \tag{29}
\]

The matrix \( \Lambda \Lambda^{T} \) is invertible if both \( q \neq 1/(4\gamma/3) \) and \( q \neq 1/(2(1-\gamma/3)) \) because in this case the rows of \( \Lambda \) are linearly independent. This is due to the fact that under the two conditions above the rows of \( \Pi^{m} \) are linearly independent, since its eigenvalues are all the possible products of three eigenvalues of \( \Pi^{m} \) [25] and the conditions above guarantee that these are nonzero. A sufficient condition for the invertibility of \( \Lambda \Lambda^{T} \) is \( q < 1/2 \), which spans most cases of interest.

Since the optimisation problem has been linearised, \( \tilde{\pi} \) may contain negative values, but in practice these are relatively small. Setting these values to zero and normalising \( \tilde{\pi} \), an approximation to the optimum distribution \( p(\text{u}|\xi') \) is obtained. An example of this approximation compared to the results of the Blahut-Arimoto algorithm is shown in Figure 15, where the high mutation rate \( q = 10^{-2} \) is used for visualisation purposes.

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